# Species Tree Inference using SVDquartets

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January 31, 2017

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## Goal of this work:

Develop a full data approach that is computationally feasible for large-scale data

## How?

- Summarize data differently, so that model requires less computation
- Develop theory to infer relationships among quartets of taxa very accurately
- Use a quartet assembly method to build a large tree

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Species tree inference using site pattern frequencies

- Data: DNA sequences for gene *i*, *D<sub>i</sub>*
- Example:

Taxon	Sequence		
(A) Human	GCCG	A	TGCCGATGCCGAA
(B) Chimp	GCCG	T	TGCCGTTGCCGTT
(C) Gorilla	GCGG	A	AGCGGAAGCGGAA

- Assume each site in the sequence evolves independently of other sites
- Data are assumed to be an iid sample of sites:
   (D<sub>i</sub>)<sub>j</sub> = data at the tips of the tree for site j in gene i
- Consider site pattern probabilities for example,  $\tilde{p}_{ATA}$

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Species tree inference using site pattern frequencies

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Model: Species tree \rightarrow gene trees \rightarrow data
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- species tree  $\rightarrow$  gene tree ::: coalescent process
- $\bullet~gene~tree \rightarrow data$  ::: nucleotide substitution models: GTR+I+ $\Gamma$  and submodels

Idea: compute site pattern probabilities under this model for 4 taxa by enumerating all possibilities for simple models

- Tedious, but not difficult
- Look for algebraic structure in the site pattern probabilities

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Taxon	Sequence
1	ACCAATGCCGATGCCAAA
2	ACCATTGCCGATGCCATA
3	ACGAAAGCGGAAGCGAAA
4	ATGAAAGCGGAAGCCAAA

$$Flat_{12|34}(P) = \begin{pmatrix} [AA] & [AC] & [AG] & [AT] & [CA] & \cdots \\ [AA] & P_{AAAA} & P_{AAAC} & P_{AAAG} & P_{AAAT} & P_{AACA} & \cdots \\ [AC] & P_{ACAA} & P_{ACAC} & P_{ACAG} & P_{ACAT} & P_{ACCA} & \cdots \\ [AG] & P_{AGAA} & P_{AGAC} & P_{AGAG} & P_{AGAT} & P_{AGCA} & \cdots \\ [AT] & P_{ATAA} & P_{ATAC} & P_{ATAG} & P_{ATAT} & P_{ATCA} & \cdots \\ [CA] & P_{CAAA} & P_{CAAC} & P_{CAAG} & P_{CAAT} & P_{CACA} & \cdots \\ [\cdots] & \cdots & \cdots & \cdots & \cdots & \cdots \end{pmatrix}$$



Taxon	Sequence
1	ACCAATGCCGATGCCAAA
2	ACCATTGCCGATGCCATA
3	ACGAAAGCGGAAGCGAAA
4	ATGAAAGCGGAAGCCAAA

$$Flat_{12|34}(P) = \begin{pmatrix} [AA] & [AC] & [AG] & [AT] & [CA] & \cdots \\ [AA] & \mathbf{5} & P_{AAAC} & P_{AAAG} & P_{AAAT} & P_{AACA} & \cdots \\ [AC] & P_{ACAA} & P_{ACAC} & P_{ACAG} & P_{ACAT} & P_{ACCA} & \cdots \\ [AG] & P_{AGAA} & P_{AGAC} & P_{AGAG} & P_{AGAT} & P_{AGCA} & \cdots \\ [AT] & P_{ATAA} & P_{ATAC} & P_{ATAG} & P_{ATAT} & P_{ATCA} & \cdots \\ [CA] & P_{CAAA} & P_{CAAC} & P_{CAAG} & P_{CAAT} & P_{CACA} & \cdots \\ [\cdots] & \cdots & \cdots & \cdots & \cdots & \cdots \end{pmatrix}$$



Taxon	Sequence
1	ACCAATGCCGGAGCCCAAA
2	ACCATTGACGGAGCCAATA
3	ACGAAAGACGGAAGCAAAA
4	ATGAAAGTCGGAAGCTAAA

$$Flat_{12|34}(P) = \begin{pmatrix} [AA] & [AC] & [AG] & [AT] & [CA] & \cdots \\ [AA] & \mathbf{5} & p_{AAAC} & p_{AAAG} & p_{AAAT} & p_{AACA} & \cdots \\ [AC] & p_{ACAA} & p_{ACAC} & p_{ACAG} & p_{ACAT} & p_{ACCA} & \cdots \\ [AG] & p_{AGAA} & p_{AGAC} & p_{AGAG} & p_{AGAT} & p_{AGCA} & \cdots \\ [AT] & p_{ATAA} & p_{ATAC} & p_{ATAG} & p_{ATAT} & p_{ATCA} & \cdots \\ [CA] & p_{CAAA} & p_{CAAC} & p_{CAAG} & \mathbf{2} & p_{CACA} & \cdots \\ [\cdots] & \cdots & \cdots & \cdots & \cdots & \cdots & \cdots \end{pmatrix}$$



Taxon	Sequence
1	ACCAATGCCGGAGCCCAAA
2	ACCATTGACGGAGCCAATA
3	ACGAAAGACGGAAGCAAAA
4	<b>ATGAAAGTCGGAAGCTAAA</b>

$$Flat_{12|34}(P) = \begin{pmatrix} [AA] & [AC] & [AG] & [AT] & [CA] & \cdots \\ [AA] & 5 & p_{AAAC} & p_{AAAG} & p_{AAAT} & p_{AACA} & \cdots \\ [AC] & p_{ACAA} & p_{ACAC} & p_{ACAG} & p_{ACAT} & p_{ACCA} & \cdots \\ [AG] & p_{AGAA} & p_{AGAC} & p_{AGAG} & p_{AGAT} & p_{AGCA} & \cdots \\ [AT] & p_{ATAA} & p_{ATAC} & p_{ATAG} & p_{ATAT} & p_{ATCA} & \cdots \\ [CA] & p_{CAAA} & p_{CAAC} & p_{CAAG} & 2 & p_{CACA} & \cdots \\ [\cdots] & \cdots & \cdots & \cdots & \cdots & \cdots & \cdots \end{pmatrix}$$

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Taxon	Sequence
1	ACCAATGCCGGAGCCCAAA
2	ACCATTGACGGAGCCAATA
3	ACGAAAGACGGAAGCAAAA
4	<b>ATGAAAGTCGGAAGCTAAA</b>

$$Flat_{12|34}(P) = \begin{pmatrix} [AA] & [AC] & [AG] & [AT] & [CA] & \cdots \\ [AA] & 5 & p_{AAAC} & p_{AAAG} & p_{AAAT} & p_{AACA} & \cdots \\ [AC] & p_{ACAA} & p_{ACAC} & p_{ACAG} & p_{ACAT} & p_{ACCA} & \cdots \\ [AG] & p_{AGAA} & p_{AGAC} & p_{AGAG} & p_{AGAT} & p_{AGCA} & \cdots \\ [AT] & p_{ATAA} & p_{ATAC} & p_{ATAG} & p_{ATAT} & p_{ATCA} & \cdots \\ [CA] & p_{CAAA} & p_{CAAC} & p_{CAAG} & 2 & p_{CACA} & \cdots \\ [\cdots] & \cdots & \cdots & \cdots & \cdots & \cdots & \cdots \end{pmatrix}$$

# These two columns are identical - matrix rank is reduced by one

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#### Results

# Main Result (Chifman and Kubatko, 2015):

- Species tree inference: For a flattening matrix constructed on the true four-taxon tree, **the matrix rank is 10** under the following model
  - species tree  $\rightarrow$  gene tree ::: coalescent process
  - ▶ gene tree  $\rightarrow$  data ::: nucleotide substitution models: GTR+I+ $\Gamma$  and submodels

# New Result (Long and Kubatko, 2017):

• This result holds even in the absence of a molecular clock or when population sizes change along the tree

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#### What about the incorrect tree?



Taxon	Sequence
1	ACCAATGCCGGAGCCCAAA
2	ACCATTGACGGAGCCAATA
3	<b>ACGAAAGACGGAAGCAAAA</b>
4	ATGAAAGTCGGAAGCTAAA

$$\mathsf{Flat}_{13|24}(\mathsf{P}) = \begin{pmatrix} [AA] & [\mathsf{AC}] & [AG] & [AT] & [\mathsf{CA}] & \cdots \\ [AA] & \mathsf{5} & \mathsf{PAAAC} & \mathsf{PAAAG} & \mathsf{PAAAT} & \mathsf{PAACA} & \cdots \\ [AC] & \mathsf{PACAA} & \mathsf{PACAC} & \mathsf{PACAG} & \mathsf{PACAT} & \mathsf{PACCA} & \cdots \\ [AG] & \mathsf{PAGAA} & \mathsf{PAGAC} & \mathsf{PAGAG} & \mathsf{PAGAT} & \mathsf{PAGCA} & \cdots \\ [AT] & \mathsf{PATAA} & \mathsf{PATAC} & \mathsf{PATAG} & \mathsf{PATAT} & \mathsf{PATCA} & \cdots \\ [CA] & \mathsf{PCAAA} & \mathsf{PCAAC} & \mathsf{PCAAG} & \mathsf{2} & \mathsf{PCACA} & \cdots \\ [\cdots] & \cdots & \cdots & \cdots & \cdots & \cdots & \cdots \end{pmatrix}$$

• Arbitrary number of states,  $\kappa$ , under the coalescent model:

- If A|B is a valid split for a tree T, then  $rank(Flat_{A|B}(P)) \leq {\binom{\kappa+1}{2}}$ .
- If C|D is not a valid split for a tree T, then  $rank(Flat_{C|D}(P)) > \binom{\kappa+1}{2}$ .
- The species tree is completely determined by knowledge of valid splits on all quartets.
- Single underlying gene tree (no coalescent assumption):
  - If A|B is a valid split for a tree T, then  $rank(Flat_{A|B}(P)) \leq 4$ .
  - If C|D is not a valid split for a tree T, then  $rank(Flat_{C|D}(P)) = 16$ .
  - The species tree is completely determined by knowledge of valid splits on all quartets.

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# • Basic idea:

- Data: aligned DNA sequences for multiple loci or for a collection of SNPs
- Construct the flattening matrix
- Compute some measure of how close the observed flattening matrix is to a matrix with rank 10

We use singular value decomposition (SVD) of the flattening matrix – define the SVD score for a split A|B to be

$$SVDscore(Flat_{A|B}(\hat{P})) = \sqrt{\sum_{i=11}^{16} \sigma_i^2}$$

where  $\sigma_i^2$  is the *i*<sup>th</sup> singular value of the matrix  $Flat_{A|B}(\hat{P})$ .

 Pick tree relationships that give the best value of the measure in the previous step

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#### Multi-locus vs. SNP data

The theory is developed for the SNP setting – why do we think this might be ok for multilocus data?

Consider the case of three possible gene trees with the probabilities below under the coalescent model:

- Gene tree  $1 p_1 = 0.4$
- Gene tree 2  $p_2 = 0.3$
- Gene tree  $3 p_3 = 0.3$

Now suppose we observe multilocus data for 1,000 genes as follows:

- Gene tree1 380 genes
- Gene tree 2 300 genes
- Gene tree 3 320 genes

Then, if the genes are equal in length, the proportion of sites coming from each tree is approximately what is predicted under the SNP model.

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Main idea: use the observed site pattern distribution to provide information about which of the three possible splits for a set of four taxa is the true split.



Compute a score for each split in a given quartet of taxa and choose the split with the best (lowest) score.

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Simulation study 1 - can we detect the correct split?

Simulate data from the Jukes-Cantor model for a 4-taxon tree and examine split scores First row: 5,000 SNP sites; Second row: 10 genes of 500bp



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Simulation study 1 - can we detect the correct split?

Simulate data from the  $GTR+I+\Gamma$  model for a 4-taxon tree and examine split scores First row: 5,000 SNP sites; Second row: 10 genes of 500bp



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#### Simulation study 1 - can we detect the correct split?

Change in scores as amount of data increases



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How do we assess variability?

- How can we measure confidence in the inferred split?
- Use a nonparametric bootstrap procedure
  - Generate bootstrap data sets from the original data matrix
  - Compute split scores on all three splits for each bootstrap data matrix
  - Record the number of bootstrap data sets for which each split is inferred, and use the proportion of these as a bootstrap support measure
- Evaluate performance of the bootstrap procedure using the same simulated data

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#### Assessing support using the bootstrap

Simulate data from the Jukes-Cantor model for a 4-taxon tree and examine bootstrap support scores



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#### Assessing support using the bootstrap

Simulate data from the  $\mathsf{GTR}{+}\mathsf{I}{+}\Gamma$  model for a 4-taxon tree and examine bootstrap support scores



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#### Extension to larger trees

# Algorithm

- Generate all quartets (small problems) or sample quartets (large problems)
- Stimate the correct quartet relationship for each sampled quartet
- Use a quartet assembly method to build the tree PAUP\* uses the method of Reaz-Bayzid-Rahman (2014), called QFM, to build the tree.





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- Multiple lineages are handled as follows:
  - Sample four species
  - Select one lineage at random from each species
  - **③** Estimate the quartet relationships among the four sampled lineages
  - Restore the species labels (but lineage quartets are saved, too)

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Simulation under more realistic scenarios Consider the effects of:

- Larger trees: 10 species
- Multi-locus data: 10-100 genes
- Varying levels of ILS: speciation intervals of 0.5, 1.0, and 2.0
- Lineage-specific rate variation: modeled by the lognormal distribution



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Simulation study 2 – larger trees with lineage-specific rate variation Average (over 100 reps) scaled RF distance (range 0 - 1)



black = lognormal(m = -0.005, s = 0.1)red = lognormal(m = -0.125, s = 0.5) blue = lognormal(m = -0.5, s = 1.0)

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500bp per gene

	10 genes	20 genes	50 genes	100 genes
Short	0.246	0.169	0.039	0.001
(0.5)	0.290	0.161	0.043	0.004
	0.290	0.160	0.050	0.004
Medium	0.117	0.024	0.001	0
(1.0)	0.107	0.027	0	0
	0.099	0.001	0.001	0
Long	0.016	0.001	0	0
(2.0)	0.017	0	0	0
	0.011	0.001	0	0



• Data: 7 (sub)species, 26 individuals (52 sequences), 19 genes

Species	Location	No. of individuals per gene
S. catenatus catenatus	Eastern U.S. and Canada	9
S. c. edwardsii	Western U.S.	4
S. c. tergeminus	Western and Central U.S.	5
S. miliarius miliarius	Southeastern U.S.	1
S. m. barbouri	Southeastern U.S.	3
S. m. streckerii	Southeastern U.S.	2
Agkistrodon sp. (outgroup)	U.S.	2

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# Empirical example: Sistrurus rattlesnakes All quartets and 100 bootstrap replicates $\sim 11 \text{ minutes}$



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# **Empirical example: Sistrurus rattlesnakes** All quartets and 100 bootstrap replicates $\sim 11 \text{ minutes}$



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#### New features: branch length estimation

- Consider the JC69 model for the symmetric species tree with 4 taxa
- In this case, there are 9 distinct site pattern probabilities
- Chifman and Kubatko (2015) showed that these site pattern probabilities under the coalescent model could be expressed as



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$$p_{i_{a}i_{b}i_{c}i_{d}|(S,\tau)} = c_{0} + c_{1}x_{1}^{2\mu} + c_{2}x_{2}^{2\mu} + c_{3}x_{1}^{2\mu}x_{2}^{2\mu} + c_{4}x_{3}^{2\mu} + c_{5}x_{1}^{\mu}x_{3}^{2\mu} + c_{6}x_{2}^{\mu}x_{3}^{2\mu} + c_{7}x_{1}^{\mu}x_{2}^{\mu}x_{3}^{2\mu} + c_{8}x_{1}^{-\frac{2}{\theta}}x_{2}^{-\frac{2}{\theta}}x_{3}^{4(\mu+\frac{1}{\theta})}$$

where  $x_j = e^{-\tau_j}$  for j = 1, 2, 3 and the coefficients are functions of the mutation rate  $\mu$  and effective population size  $\theta$ .

#### New features: branch length estimation

• Let  $C_{9 \times 9}$  be the matrix of coefficients. Then the above expressions for the site pattern probabilities can be written as

$$\mathbf{C}\beta = p$$

where

$$\beta' = \begin{pmatrix} x_1^{2\mu}, x_2^{2\mu}, x_1^{2\mu}x_2^{2\mu}, x_3^{2\mu}, x_1^{\mu}x_3^{2\mu}, x_2^{\mu}x_3^{2\mu}, x_1^{\mu}x_2^{2\mu}x_3^{2\mu}, (x_1x_2)^{-2/\theta}x_3^{4(\mu+1/\theta)} \end{pmatrix}$$
  
and  
$$p' = (p_{\text{xxxx}}, p_{\text{xxxy}}, p_{\text{xyxx}}, p_{\text{xyxy}}, p_{\text{xxyy}}, p_{\text{xxyz}}, p_{\text{xyzx}}, p_{\text{xyzx}}, p_{\text{xyzy}})$$

- Use this to write the likelihood for the four-taxon case, and find maximum likelihood estimates numerically
- Asymptotic variances can be found using standard statistical theory (Fisher information matrix, etc.)

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#### New features: branch length estimation

- For 4 taxa:
  - Test robustness and possibly use models more general than JC69
- For larger trees:
  - Combine estimates for 4 taxa





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## SVDquartets Summary

# • Advantages:

- Quick! And scales well to large taxon sets and next-gen sequencing data
- Easily parallelized
- Intuitive method for handling missing data
- Potential for application to other data types (codons, amino acids, etc.)

# • Disadvantages:

 Estimating a matrix with 256 entries .... so may not work well with limited data

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#### **SVDQuartets**

# Now on to the tutorial!

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